

# Statistics Qualifying Exam

9:00 am - 1:00 pm, Monday, May 2, 2016

1. If  $Y_1$  and  $Y_2$  have a joint distribution given by

$$f(y_1, y_2) = \begin{cases} \frac{1}{2}y_1y_2, & 0 \leq y_2 \leq y_1 \leq 2; \\ 0, & \text{elsewhere.} \end{cases}$$

- (a) Find the marginal distributions of  $Y_1$  and  $Y_2$ .
  - (b) Find the conditional distribution of  $Y_2$  given  $Y_1 = y_1$ .
  - (c) Find the values of  $E(Y_2|Y_1 = 1)$  and  $Var(Y_2|Y_1 = 1)$ .
  - (d) What is the density of  $U = Y_1 - Y_2$ ?
2. Suppose that  $X_1, \dots, X_n$  is a random sample from Uniform( $\theta, 2\theta$ ) distribution, where  $\theta > 0$  is an unknown parameter.
- (a) Find the maximum likelihood estimator (MLE) for  $\theta$ .
  - (b) Prove that the MLE is consistent.
3. Let  $X_1, \dots, X_n$  is a random sample from a distribution with pdf as  $f(x|\theta) = \theta^{-1} \exp(-x/\theta), x \geq 0, \theta > 0$ .
- (a) Find the MLE of  $P(X \leq 2)$ .
  - (b) Find the minimum variance unbiased estimator (MVUE) of  $P(X \leq 2)$ .
  - (c) Derive the exact likelihood ratio test for the following hypothesis

$$H_0 : \theta = \theta_0 \text{ versus } H_1 : \theta \neq \theta_0.$$

Please carefully specify the likelihood ratio  $\Lambda$ , the test statistic, the sampling distribution of the test statistic under the null hypothesis, and the decision rules of a size  $\alpha$  test.

- (d) For the decision rule derived in Part (b), obtain the distribution of the test statistic under a general alternative (that is  $H_1 : \theta \neq \theta_0$ ) and use it to obtain the power function of the test.
4. Let  $X_1, \dots, X_n$  be a random sample from the pdf  $f(x|\theta) = \exp[-(x - \theta)]$ , where  $-\infty < \theta < \infty$  and  $x \geq \theta$ .
- (a) Find the minimal sufficient statistic for  $\theta$ .
  - (b) Let  $Y_1 < Y_2 < \dots < Y_n$  be the ordered sample, and define  $R_i = Y_n - Y_i, i = 1, \dots, n - 1$ . Show that the set  $(R_1, R_2, \dots, R_{n-1})$  is ancillary for  $\theta$ .
  - (c) Assume it is given that  $Y_1 = \min_i X_i$  is a complete sufficient statistic. Show that  $Y_1$  and  $S^2 = \sum_{i=1}^n (X_i - \bar{X})^2 / (n - 1)$  are independent, where  $\bar{X} = \sum_{i=1}^n X_i / n$ .
  - (d) Now prove the condition given in Part (c). That is, show that  $Y_1 = \min_i X_i$  is a complete sufficient statistic.

5. The following is part of ANOVA table for a simple linear regression model

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i,$$

where  $\varepsilon_i$ 's are i.i.d. from  $N(0, \sigma^2)$ ,  $i = 1, \dots, n$ , and  $n$  is the number of observations. (No need to complete the table.)

| Source          | DF  | Sum of Squares | Mean Square | F Value | Pr > F |
|-----------------|-----|----------------|-------------|---------|--------|
| Model           | *** | *****          | 252378      | 105.88  | <.0001 |
| Error           | 23  | *****          | *****       |         |        |
| Corrected Total | *** | *****          |             |         |        |

- (a) Compute the coefficient of determination  $R^2$ .
- (b) Assume that we now know the least square estimate of  $\beta_1$  is  $b_1 = 3.57$ . Construct a two-sided  $t$ -test of whether or not  $\beta_1 = 3$ . State the null and alternative hypotheses, the value of the test statistic, the sampling distribution under the null hypothesis and the decision rule.
6. A gardener is interested in studying the relationship between fertilizer and tomato yield. The gardener has two gardens (1 and 2). He divides each into 9 plots. Three fertilizer application rates (3, 5, and 7 units/acre) are assigned to the plots in garden 1 in a completely randomized fashion. The same three fertilizer application rates (3, 5, and 7 units/acre) are assigned to the plots in garden 2 in a completely randomized fashion. Thus there are three plots for each combination of garden and fertilizer application rate. After some initial analyses, the gardener decides to base his analysis on the following SAS code and output.

- (a) Note that rate was not included in the class statement. What would the Model and Error DF change to if rate were included in the class statement? That is, complete the following tables by filling in the missing values for DF (*you only need to provide DF for ???, but do not need to calculate any of the Sum Squares*).

| Source          | DF  | Sum of Sum Squares |
|-----------------|-----|--------------------|
| Model           | ??? |                    |
| Error           | ??? |                    |
| Corrected Total | ??? |                    |

| Source      | DF  | Type I SS |
|-------------|-----|-----------|
| garden      | ??? |           |
| rate        | ??? |           |
| rate*garden | ??? |           |

- (b) Estimate the equation of the regression line relating yield to fertilizer application rate in garden 1.
- (c) Estimate the equation of the regression line relating yield to fertilizer application rate in garden 2.
- (d) Is there a significant difference between the slopes of the two regression lines? To get full credits, give an appropriate test statistic,  $p$ -value, and conclusion, using  $\alpha = 0.05$ .
- (e) Suppose the gardener were to apply 7 units of fertilizer per acre to all plots in both gardens. Which garden would have the higher expected yield?

```

proc glm;
  class garden;
  model yield=garden rate garden*rate / solution;
run;

```

The GLM Procedure  
 Dependent Variable: yield

| Source          | DF | Sum of Squares | Mean Square | F Value | Pr > F |
|-----------------|----|----------------|-------------|---------|--------|
| Model           | 3  | 58.88888889    | 19.62962963 | 27.33   | <.0001 |
| Error           | 14 | 10.05555556    | 0.71825397  |         |        |
| Corrected Total | 17 | 68.94444444    |             |         |        |

| Source      | DF | Type I SS   | Mean Square | F Value | Pr > F |
|-------------|----|-------------|-------------|---------|--------|
| garden      | 1  | 2.72222222  | 2.72222222  | 3.79    | 0.0719 |
| rate        | 1  | 52.08333333 | 52.08333333 | 72.51   | <.0001 |
| rate*garden | 1  | 4.08333333  | 4.08333333  | 5.69    | 0.0318 |

| Parameter     | Estimate | Standard Error | t Value | Pr >  t |
|---------------|----------|----------------|---------|---------|
| Intercept     | -1.11 B  | 0.90993803     | -1.22   | 0.2422  |
| garden 1      | 3.69 B   | 1.28684670     | 2.87    | 0.0123  |
| garden 2      | 0.00 B   | .              | .       | .       |
| rate          | 1.33 B   | 0.17299494     | 7.71    | <.0001  |
| rate*garden 1 | -0.58 B  | 0.24465179     | -2.38   | 0.0318  |
| rate*garden 2 | 0.00 B   | .              | .       | .       |

7. An experiment is conducted to study the effects of loading frequencies (Frequency) and environmental conditions (Environment) on fatigue crack growth at a constant 22 MPa stress for a particular material. The data from this experiment are shown below (the response is crack growth rate):

| Frequency | Environment      |                     |                         |
|-----------|------------------|---------------------|-------------------------|
|           | 1: Air           | 2: H <sub>2</sub> O | 3: SaltH <sub>2</sub> O |
| 1         | 2.29, 2.47, 2.12 | 2.86, 3.03, 2.73    | 4.93, 4.75, 5.06        |
| 2         | 3.15, 2.88, 2.56 | 4.00, 4.44, 4.70    | 3.10, 3.24, 3.98        |
| 3         | 2.24, 2.71, 2.81 | 4.00, 4.30, 3.20    | 4.86, 4.26, 5.20        |

Some summary statistics are give below.

|                   |       |             |       |
|-------------------|-------|-------------|-------|
| grand mean: 3.551 |       |             |       |
| frequency         | MEAN  | environment | MEAN  |
| 1                 | 3.360 | 1           | 2.581 |
| 2                 | 3.561 | 2           | 3.696 |
| 3                 | 3.731 | 3           | 4.376 |

|           |             |       |
|-----------|-------------|-------|
| frequency | environment | MEAN  |
| 1         | 1           | 2.293 |
| 1         | 2           | 2.873 |
| 1         | 3           | 4.913 |
| 2         | 1           | 2.863 |
| 2         | 2           | 4.380 |
| 2         | 3           | 3.440 |
| 3         | 1           | 2.587 |
| 3         | 2           | 3.833 |
| 3         | 3           | 4.773 |

Suppose the following statistical model is used to fit the data.

$$Y_{ijk} = \mu + \tau_i + \beta_j + (\tau\beta)_{ij} + \epsilon_{ijk}; k = 1, 2, 3$$

where  $\tau_i (i = 1, 2, 3)$  and  $\beta_j (j = 1, 2, 3)$  are the main effects of frequency, the main effects of environment, respectively, and  $(\tau\beta)_{ij}$  are their interactions. For parameter estimation, we impose the following constraints:  $\sum_i \tau_i = \sum_j \beta_j = \sum_i (\tau\beta)_{ij} = \sum_j (\tau\beta)_{ij} = 0$ .

- Calculate the estimates of  $\tau_1, \tau_2, \tau_3$  and  $(\tau\beta)_{23}$ .
- Calculate the sum of squares due to frequency.

- (c) The ANOVA of the data was done in SAS and the output is shown. Notice that some quantities are removed. Test if the main effects of frequency are significant and if the interactions of frequency and environment are significant (using  $\alpha = 5\%$ ).

Dependent Variable: crack

| Source     | DF | Sum of Squares | Mean Square | F Value | Pr > F |
|------------|----|----------------|-------------|---------|--------|
| Model      | 8  | 22.719         | 2.840       | 22.11   | <.0001 |
| Error      | 18 | 2.312          | 0.128       |         |        |
| Coed Total | 26 | 25.031         |             |         |        |

| Source                | DF | Type I SS | Mean Square | F Value | Pr > F |
|-----------------------|----|-----------|-------------|---------|--------|
| frequency             | *  | *****     | *****       | ****    | *****  |
| environment           | 2  | 14.773    | 7.387       | 57.51   | <.0001 |
| frequency*environment | *  | *****     | *****       | ****    | *****  |

- (d) Calculate the critical difference (CD) for the treatment pairwise comparison using Tukey's method. Report the results for the following (*and only the following*) pairs: (1, 1) versus (1,3); (2, 2) versus (3, 2).
8. An experiment was run to determine whether four specific temperatures affect the production of a certain type of chemical compound. The experiment led to the following data. The response is the amount of the chemical compound produced over a period of time.

| Temperature | Response |       |       |       |       | Mean $\bar{Y}_i$ | St.D. $s_i$ |
|-------------|----------|-------|-------|-------|-------|------------------|-------------|
| 100         | 8.71     | 10.47 | 9.62  | 11.55 | 10.25 | 10.12            | 1.05        |
| 125         | 24.12    | 25.23 | 26.08 | 20.30 | 21.15 | 23.38            | 2.54        |
| 150         | 30.37    | 31.14 | 35.62 | 30.19 | 27.03 | 30.87            | 3.09        |
| 175         | 28.24    | 24.15 | 22.02 | 26.46 | 27.44 | 25.66            | 2.55        |

- (a) Generate (approximately) the plot of log St.D. ( $\log s_i$ ) versus log Mean ( $\log \bar{Y}_i$ ) by hand.
- (b) Based on your plot, is the constant variance assumption for ANOVA valid? What remedy you can recommend? Derive the remedy explicitly.

The modified data are given below.

| Temperature | Transformed Response |      |      |      |      | Mean $\bar{Y}_i$ | St.D. $s_i$ |
|-------------|----------------------|------|------|------|------|------------------|-------------|
| 100         | 2.16                 | 2.35 | 2.26 | 2.45 | 2.33 | 2.310            | 0.1079      |
| 125         | 3.18                 | 3.23 | 3.26 | 3.01 | 3.05 | 3.146            | 0.1106      |
| 150         | 3.41                 | 3.44 | 3.57 | 3.41 | 3.30 | 3.426            | 0.0966      |
| 175         | 3.34                 | 3.18 | 3.09 | 3.28 | 3.31 | 3.240            | 0.1032      |

- (c) It is known that  $SST = 3.839$ . Choose an appropriate model, construct the ANOVA table, and test if the temperatures have different effects on the transformed response.

- (d) What is the estimate for  $\tau_1$  (the treatment effect at temperature 100) under the constraint  $\sum_i \tau_i = 0$ ?
- (e) Let  $\mu_1, \mu_2, \mu_3$  and  $\mu_4$  be the treatment means at temperatures 100, 125, 150 and 175, respectively, and let  $L = \mu_1 - 2\mu_2 + \mu_3$ . What is the estimate for  $L$ ?
- (f) Test  $H_0 : L = 0$  vs  $H_1 : L \neq 0$ .
- (g) One decides to use the contrasts based on orthogonal polynomials to further model the relationship between temperature and the transformed response. Part of the SAS code and output is given below. Fill in the DF, Contrast SS, Mean Square and F Value for the linear contrast.

```
contrast 'linear'      temperature -3 -1  1 3;
contrast 'quadratic'  temperature  1 -1 -1 1;
contrast 'cubic'      temperature -1  3 -3 1;
```

```
-----
Contrast   DF  Contrast SS      Mean Square    F Value    Pr > F
linear     *  *****          *****          *****    <.0001
quadratic  1  1.30560500      1.30560500     119.07     <.0001
cubic      1  0.00202500      0.00202500      0.18      0.6731
```